Probabilistic Method and Random Graphs Lecture 6. Poisson Approximation with Application ¹

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¹The slides are mainly based on Chapter 5 of the textbook *Probability and Computing* and Lecture 12 of Ryan O'Donnell's lecture notes of *Probability and Computing*.

Questions, comments, or suggestions?

A recap of Lecture 5

Poisson Convergence (LSN)

Assume that $X_n \sim Bin(n, p_n)$ with $\lim_{n\to\infty} np_n = \lambda$. For any fixed k, $\lim_{n\to\infty} \Pr(X_n = k) = \frac{e^{-\lambda}\lambda^k}{k!}$. Valid when weakly dependent.

Almost Valid when strongly dependent (Stein-Chen Theorem).

Joint distribution of bin loads

Max load:
$$L(n,n) < 3 \frac{\ln n}{\ln \ln n}$$
 with high probability.
 $\Pr(X_1 = k_1, ..., X_n = k_n) = \frac{m!}{k_1! k_2! \cdots k_n! n^m}$

Poisson approximation theorem

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$$(X_1^{(m)}, X_2^{(m)}, ... X_n^{(m)}) \sim (Y_1^{(\mu)}, Y_2^{(\mu)}, ... Y_n^{(\mu)} | \sum Y_i^{(\mu)} = m)$$

Application: coupon collector's problem

$$\lim_{n \to \infty} \Pr(X > n \ln n + cn) = 1 - e^{-e^{-c}}$$

Hard to use due to conditioning.

Can we remove the condition?

Notation

$$X_i^{(m)}$$
: the load of bin *i* in (m, n) -model.

 $Y_i^{(m)}$: independent Poisson r.v.s with expectation $\frac{m}{n}$.

Theorem

For any non-negative $n\text{-}\mathrm{ary}$ function f, we have $\mathbb{E}[f(X_1^{(m)},...X_n^{(m)})] \leq e\sqrt{m}\mathbb{E}[f(Y_1^{(m)},...Y_n^{(m)})].$

Remark

Unlike $(X_1^{(m)}, X_2^{(m)}, ...X_n^{(m)}) \sim (Y_1^{(\mu)}, Y_2^{(\mu)}, ...Y_n^{(\mu)}| \sum Y_i^{(\mu)} = m)$, the mean of the Poisson distribution is $\frac{m}{n}$, not arbitrary. Condition-freedom at the cost of approximation.

Proof

$$\begin{split} \mathbb{E}[f(Y_1^{(m)}, \dots Y_n^{(m)})] \\ &= \sum_k \mathbb{E}[f(Y_1^{(m)}, \dots Y_n^{(m)})| \sum_i Y_i^{(m)} = k] \Pr(\sum_i Y_i^{(m)} = k) \\ &\geq \mathbb{E}[f(Y_1^{(m)}, \dots Y_n^{(m)})| \sum_i Y_i^{(m)} = m] \Pr(\sum_i Y_i^{(m)} = m) \\ &= \mathbb{E}[f(X_1^{(m)}, \dots X_n^{(m)})] \Pr(\sum_i Y_i^{(m)} = m). \end{split}$$

$$\begin{split} \sum_i Y_i^{(m)} \sim Poi(m) \Rightarrow \Pr(\sum_i Y_i^{(m)} = m) = \frac{m^m e^{-m}}{m!} \geq \frac{1}{e\sqrt{m}} \text{ since } \\ m! < e\sqrt{m}(me^{-1})^m. \end{split}$$

Remark

$$\mathbb{E}[f(X_1^{(m)},...X_n^{(m)})] \leq 2\mathbb{E}[f(Y_1^{(m)},...Y_n^{(m)})]$$
 if f is monotonic in m

প্র 6/14 Any event that takes place with probability p in the independent Poisson coupling takes places in Bins&Balls setting with probability at most $pe\sqrt{m}$

If the probability of an event in Bins&Balls is monotonic in m, it is at most twice of that in the independent Poisson coupling

Remark

Use approximate experiments to bound the exact-case probability. Powerful in bounding the probability of rare events in Bins&Balls.

Application

Lower bound of max load in (n, n)-model

Asymptotically, $\Pr(\mathcal{E}) \leq \frac{1}{n}$, where \mathcal{E} is the event that the max. load in the (n, n)-Bins&Balls model is smaller than $\frac{\ln n}{\ln \ln n}$.

Remark: In fact, the max. load is $\Theta\left(\frac{\ln n}{\ln \ln n}\right)$ w.h.p.

Proof

In the Poisson approximation, a bin has at least $M = \frac{\ln n}{\ln \ln n}$ balls with probability at least $\frac{1}{eM!} \Rightarrow \Pr(\mathcal{E}') \le \left(1 - \frac{1}{eM!}\right)^n \le e^{-\frac{n}{eM!}}$.

$$M! \le e\sqrt{M}(e^{-1}M)^M \le M(e^{-1}M)^M$$

$$\Rightarrow \ln M! \le \ln n - \ln \ln n - \ln(2e) \Rightarrow M! \le \frac{n}{2e\ln n}.$$

Altogether, $\Pr(\mathcal{E}) \leq e\sqrt{n} \Pr(\mathcal{E}') \leq \frac{e\sqrt{n}}{n^2} \leq \frac{1}{n}$.

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Application: Hashing

Used to look up records, protect data, find duplications ...

Membership problem: password checker

Binary search vs Hashing

Hash table (1953, H. P. Luhn @IBM)

Hash functions: efficient, deterministic, uniform, non-invertible Random: coin tossing, SUHA SHA-1 (broken by Wang et al., 2005) Bins&Balls model

Efficiency

Search time for m words in n bins: expected vs worst. Space: $\geq 256m$ bits if each word has 256 bits. Potential wasted space: $\frac{1}{e}$ in the case of m = n. Trade space for time. Can we improve space-efficiency?

Fingerprint

Succinct identification of lengthy information

Fingerprint hashing

Fingerprinting \rightsquigarrow sorting fingerprints (rather than original data) \rightsquigarrow binary search.

Trade time for space

Performance

False positive: due to loss of information No other errors Partial correction using white lists

Probability of a false positive: m words, b bits

Fingerprint of an acceptable differs from that of a bad: $1 - \frac{1}{2^b}$. Probability of a false positive: $1 - \left(1 - \frac{1}{2^b}\right)^m \ge 1 - e^{-\frac{m}{2^b}}$.

Determine b

For a constant c, false positive $< c \Rightarrow e^{-\frac{m}{2b}} \ge 1 - c$. So, $b \ge \log_2 \frac{-m}{\ln(1-c)} = \Omega(\ln m)$. If $b \ge 2\log_2 m$, false positive $< \frac{1}{m}$. 2^{16} words, 32-bit fingerprints, false positive $< 2^{-16}$. Save a factor of 8 if each word has 256 bits.

Can more space be saved while getting more time-efficient?

Bloom Filter

1970, CACM, by Burton H. Bloom.

Used in Bigtable and HBase.

Basic idea

Hash table + fingerprinting Illustration

False positive is the only source of errors.

False positive: m words, n-bit array, k mappings

A specific bit is 0 with probability $(1 - \frac{1}{n})^{km} \approx e^{-\frac{km}{n}} \triangleq p$. Resonable to assume that a fraction p of bits are 0. By Poisson approximation and Chernoff bounds. False positive probability: $f \triangleq (1 - (1 - \frac{1}{n})^{km})^k \approx (1 - e^{-\frac{km}{n}})^k$

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Determine k for fixed m, n

Objective

Minimize f. Dilemma of k: chances to find a 0-bit vs the fraction of 0-bits.

Optimal k

$$\begin{split} \frac{d \ln f}{dk} &= \ln \left(1 - e^{-\frac{km}{n}} \right) + \frac{km}{n} \frac{e^{-\frac{km}{n}}}{1 - e^{-\frac{km}{n}}}.\\ \frac{d \ln f}{dk}|_{k=\frac{n}{m} \ln 2} &= 0.\\ f|_{k=\frac{n}{m} \ln 2} &= 2^{-k} \approx 0.6185^{n/m}.\\ f &< 0.02 \text{ if } n = 8m, \text{ and } f < 2^{-16} \text{ if } n = 23m, \text{ saving 1/4 space} \end{split}$$

Remark

Fix n/m, the #bits per item, and get a constant error probability. In fingerprint hashing, $\Omega(\ln m)$ bits per item guarantee a constant error probability

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Lecture 12 of the CMU lecture notes by Ryan O'Donnell.