

SOLUTION MANUAL FOR *PROBABILISTIC METHOD AND RANDOM GRAPHS*

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ABSTRACT. In this note, we present solutions for selected homework problems for course *Probabilistic method and random graphs*: <http://z14120902.github.io/pm.html>.

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1. HW1

1. Memoryless implies

$$\forall m, n \in \mathbb{Z}^+ : \mathbf{P}(X > m + n \mid X > m) = \mathbf{P}(X > n)$$

or

$$\forall m, n \in \mathbb{Z}^+ : \frac{\mathbf{P}(X > m + n \cap X > m)}{\mathbf{P}(X > m)} = \frac{\mathbf{P}(X > m + n)}{\mathbf{P}(X > m)} = \mathbf{P}(X > n)$$

or

$$\forall m, n \in \mathbb{Z}^+ : \mathbf{P}(X > m + n) = \mathbf{P}(X > m)\mathbf{P}(X > n)$$

the rest will be easy.

2.

1. Recall *Coupon collector's problem*, there are 2 coupons in this case

$$\mathbf{E}X = 1 + 2 = 3$$

2.

- after first kid, success within 5 kids: $\sum_{i=1}^4 \frac{i}{2^i} = \frac{13}{8}$
- after first kid, fail within 5 kids: $4 \cdot \frac{1}{2^4} = \frac{1}{4}$

$$\mathbf{E}X = 1 + \frac{13}{8} + \frac{1}{4} = \frac{23}{8}$$

2. HW2

1.

$$\mathbf{P}(X \geq (1 + \delta)\mu_H) = \mathbf{P}(e^{\lambda X} \geq e^{\lambda(1+\delta)\mu_H}) \leq \frac{\mathbf{E}[e^{\lambda X}]}{e^{\lambda(1+\delta)\mu_H}} = \frac{\mathbf{E}[e^{\lambda \sum_{i=1}^n X_i}]}{e^{\lambda(1+\delta)\mu_H}} = \frac{e^{\sum_{i=1}^n (p_i e^{\lambda} + (1-p_i))}}{e^{\lambda(1+\delta)\mu_H}} \leq \frac{e^{\sum_{i=1}^n p_i (e^{\lambda} - 1)}}{e^{\lambda(1+\delta)\mu_H}} = \frac{e^{\mu(e^{\lambda} - 1)}}{e^{\lambda(1+\delta)\mu_H}} \leq \frac{e^{\mu_H(e^{\lambda} - 1)}}{e^{\lambda(1+\delta)\mu_H}} = \left(\frac{e^{(e^{\lambda} - 1)}}{e^{\lambda(1+\delta)}} \right)^{\mu_H}$$

the rest will be easy.

2. we need the following

$$\mathbf{E}[e^{\lambda a_i X_i}] = p_i e^{\lambda a_i} + 1 - p_i = 1 + p_i (e^{\lambda a_i} - 1) \leq \frac{e^{p_i (e^{\lambda a_i} - 1)}}{e^{p_i a_i (e^{\lambda} - 1)}}$$

or

$$e^{\lambda a_i} - 1 \leq a_i (e^{\lambda} - 1)$$

or

$$\frac{e^{\lambda a_i} - 1}{a_i} \leq \frac{e^{\lambda \cdot 1} - 1}{1}$$

this is slope of line through $(x, e^{\lambda x})$ and $(0, 1)$, which is obvious via plot of function $e^{\lambda x}$, the rest will be easy.

3.

1.

$$\mathbf{E}f(Z) = \sum_i p_i f(z_i) = \sum p_i f(z_i * 1 + (1-z_i) * 0) \leq \sum_i p_i z_i f(1) + p_i (1-z_i) f(0) = pf(1) + (1-p)f(0) = \mathbf{E}f(X)$$

2. Omit

3. HW4

1. the exact probability is

$$\mathbf{P}(n \text{ bins } m \text{ balls max load} = 1) = \left(1 - \frac{1}{n}\right) \left(1 - \frac{2}{n}\right) \dots \left(1 - \frac{m-1}{n}\right)$$

1. we need $\mathbf{P}(n \text{ bins } m \text{ balls max load} = 1) \leq \frac{1}{e}$ or

$$\mathbf{P}(n \text{ bins } m \text{ balls max load} = 1) \leq e^{-\frac{1}{n}} e^{-\frac{2}{n}} \dots e^{-\frac{m-1}{n}} = e^{-\frac{m(m-1)}{2n}} \leq \frac{1}{e}$$

we can calculate m

2. we need $\mathbf{P}(n \text{ bins } m \text{ balls max load} = 1) \geq \frac{1}{2}$ or

$$\mathbf{P}(n \text{ bins } m \text{ balls max load} = 1) \geq e^{-\frac{1}{n} - \frac{1}{n^2} - \frac{2}{n} - \frac{2^2}{n^2} \dots - \frac{m-1}{n} - \frac{(m-1)^2}{n^2}} = e^{-\frac{m(m-1)}{2n} - \frac{(m-1)m(2m-1)}{6n^2}} \geq \frac{1}{2}$$

we can calculate m

2.

1. $\mathbf{P}(X = n) = e^{-\mu} \frac{\mu^n}{n!}$

$$\begin{aligned} \mathbf{P}(Y = k) &= \sum_{n=k}^{\infty} \mathbf{P}(X = n) \binom{n}{k} p^k (1-p)^{n-k} = \sum_{n=k}^{\infty} e^{-\mu} \frac{\mu^n}{n!} \binom{n}{k} p^k (1-p)^{n-k} \\ &= \frac{e^{-\mu} p^k}{k!} \sum_{n=k}^{\infty} \frac{\mu^n (1-p)^{n-k}}{(n-k)!} = \frac{e^{-\mu} (\mu p)^k}{k!} \sum_{n=k}^{\infty} \frac{\mu^{n-k} (1-p)^{n-k}}{(n-k)!} = \frac{e^{-\mu} (\mu p)^k}{k!} e^{\mu(1-p)} = e^{-\mu p} \frac{(\mu p)^k}{k!} \end{aligned}$$

Z can be proved likewise

2.

$$\mathbf{P}(Y = k_1, Z = k_2) = \mathbf{P}(X = k_1 + k_2) \binom{k_1 + k_2}{k_1} p^{k_1} (1-p)^{k_2} = e^{-\mu} \frac{\mu^{k_1 + k_2}}{(k_1 + k_2)!} \frac{(k_1 + k_2)!}{k_1! k_2!} p^{k_1} (1-p)^{k_2} = e^{-\mu p} \frac{(\mu p)^{k_1}}{k_1!} e^{-\mu(1-p)} \frac{(\mu(1-p))^{k_2}}{k_2!} = \mathbf{P}(Y = k_1) \mathbf{P}(Z = k_2)$$

3.

1. for $n = 1$, there are no other students

$$\mathbf{P}(2 \text{ students same birthday}) = 0$$

for $n \in \{2, \dots, 365\}$

the probability of max load is 1 is

$$\mathbf{P}(\text{max load} = 1) = \left(1 - \frac{1}{365}\right) \left(1 - \frac{2}{365}\right) \dots \left(1 - \frac{n-1}{365}\right)$$

thus

$$\mathbf{P}(2 \text{ students same birthday}) = 1 - \left(1 - \frac{1}{365}\right) \left(1 - \frac{2}{365}\right) \dots \left(1 - \frac{n-1}{365}\right)$$

for $n \in \{365, 366, \dots\}$

$$\mathbf{P}(2 \text{ students same birthday}) = 1$$

2. for $n = 1$, there are no other students

$$\mathbf{P}(\text{existing another students same birthday}) = 0$$

for $n \in \{2, \dots, 365\}$

$$\mathbf{P}(\text{no other students same birthday}) = \left(\frac{364}{365}\right)^{n-1}$$

$$\mathbf{P}(\text{existing another students same birthday}) = 1 - \left(\frac{364}{365}\right)^{n-1}$$

4. $\mathbf{P}(X \geq x) \leq \frac{\mathbf{E}e^{\lambda X}}{e^{\lambda x}} = \frac{e^{\mu(e^{\lambda} - 1)}}{e^{\lambda x}}$. let $\lambda = \ln\left(\frac{x}{\mu}\right)$

4. HW5

1. This is called Brun's sieve, Alon's book chapter 8.3.

2.

1. probability of a bin with load 1

$$\mathbf{P}(X_i = 1) = \binom{b}{1} \frac{1}{n} \left(1 - \frac{1}{n}\right)^{b-1}$$

the expected balls will be served

$$\mathbf{E}X = n\mathbf{P}(X_i = 1) = b \left(1 - \frac{1}{n}\right)^{b-1}$$

thus, expected number of balls at the start of the next round $b - b \left(1 - \frac{1}{n}\right)^{b-1}$

2.

$$x_{j+1} = x_j - x_j \left(1 - \frac{1}{n}\right)^{x_j-1} = x_j \left[1 - \left(1 - \frac{1}{n}\right)^{x_j-1}\right]$$

consider $f(x) = \left(1 - \frac{1}{n}\right)^{x+1} - \left(1 - \frac{x}{n}\right)$, we can get $x_{j+1} \leq \frac{x_j^2}{n}$ or $\ln x_{j+1} \leq 2 \ln x_j - \ln n$ the rest will be easy.

3. Recall *Poisson Approximation Theorem*. $\mathbf{P}(X_1 \neq 0 \cap \dots \cap X_n \neq 0 \mid \sum X_i = k)$ is the same probability of all bins are not empty in k balls into n bins model, the probability increases with k for the obvious reason.

4. Recall *Poisson Approximation Theorem*. $\lim_{n \rightarrow \infty} \mathbf{P}(\mathcal{E} \mid X = m + \sqrt{2m \ln m}) - \mathbf{P}(\mathcal{E} \mid X = m - \sqrt{2m \ln m}) \leq$ the probability of at least one empty bin after throwing $m - \sqrt{2m \ln m}$ balls but at least one among the next $2\sqrt{2m \ln m}$ balls goes into that bin

$$\leq \lim_{n \rightarrow \infty} \frac{2\sqrt{2m \ln m}}{n} = \lim_{n \rightarrow \infty} \frac{2\sqrt{2n \ln n \ln(n \ln n)}}{n} = \lim_{n \rightarrow \infty} \frac{2\sqrt{2n(\ln n)^2 + 2n \ln n \ln \ln n}}{n} \sim \lim_{n \rightarrow \infty} \frac{\ln n}{\sqrt{n}} \rightarrow 0$$

5. HW6

1.

$$1. m = n, \lambda = 1: e\sqrt{n} \left(e^{-1} \frac{1^n}{1!}\right)^n \leq \sqrt{ne} e^{1-n}$$

$$2. \frac{n!}{n^n}$$

2. [Mitzenmacher and Upfal, 2005] Theorem 5.10

6. HW7

1. for any graph G over n vertices, we need to prove for model \mathcal{G}_n and $\mathcal{G}_{n, \frac{1}{2}}$

$$\mathbf{P}_{\mathcal{G}_n}(G) = \mathbf{P}_{\mathcal{G}_{n, \frac{1}{2}}}(G) = \frac{1}{2^{\binom{n}{2}}}$$

2. [Erdos and R&WI, 1959] Theorem 1

7. HW8

1. we need

$$p = \mathbf{P}\left(|S| \geq \frac{|V|}{D+1}\right) \geq \frac{1}{2D|V|^2}$$

since

$$\mathbf{E}|S| \geq \sum_{i=1}^n \frac{1}{d_i + 1} \geq \frac{|V|}{D+1}$$

$$\mathbf{E}|S| = \sum_{|S| \geq \frac{|V|}{D+1}} |S| \mathbf{P}(|S|) + \sum_{|S| < \frac{|V|}{D+1}} |S| \mathbf{P}(|S|) \leq |V|p + \left(\frac{|V|}{D+1} - 1\right)(1-p)$$

thus

$$|V|p + \left(\frac{|V|}{D+1} - 1\right)(1-p) \geq \frac{|V|}{D+1}$$

thus

$$p \left(\frac{D|V|}{D+1} + 1\right) \geq 1$$

thus

$$p \geq \frac{D+1}{D+1+D|V|} > \frac{1}{2D|V|^2}$$

2.

1. It is an independent set because of the construction process: for each vertex i , $i \in S(\sigma)$ if and only if no neighbor j of i precedes i in the permutation σ . For vertex i and its neighbors $\Gamma(i)$, only one of them can be chosen.

2. random permutation

3.

1. consider 2-color edges of K_n

$$\mathbf{P}(\text{a 4-clique is monochromatic}) = 2 * \frac{1}{2^{\binom{4}{2}}} = 2^{-5}$$

thus

$$\text{total number of monochromatic copies of } K_4 \leq \binom{n}{4} 2^{-5}$$

2. randomly assign color to edges

8. HW9

1.

$$\mathbf{E}[\text{total number of monochromatic copies of } K_k \text{ for 2-coloring the edges of } K_n] = 2 * \binom{n}{k} \frac{1}{2^{\binom{k}{2}}} = \binom{n}{k} 2^{1-\binom{k}{2}}$$

we remove 1 vertex from each above monochromatic copies of K_k , we will have a $K_x : x = n - \binom{n}{k} 2^{1-\binom{k}{2}}$. And there are no monochromatic copies of K_k .

2. Use the derandomization technique mentioned in the class. Choose each assignment that makes total number of monochromatic copies below its expectation.

Algorithm 1 find monochromatic K_4

edges := x_1, x_2, \dots, x_m
color choices := $v_k \in \{0, 1\}$
for $k = 1$ **to** m **do**

$$x_k = \underset{v_k \in \{0,1\}}{\operatorname{argmin}} \mathbf{E}[\text{total number of monochromatic } K_4 \mid x_1 = v_1, \dots, x_{k-1} = v_{k-1}, x_k = v_k]$$

end for

3. similar idea

Algorithm 2 permutation σ

vertices $V := \{v_1, v_2, \dots, v_n\}$
permuted vertices $X := \{x_1, x_2, \dots, x_n\}$
for $k = 1$ **to** n **do**

$$x_k = \underset{v_k \in V \setminus \{v_1, v_2, \dots, v_{k-1}, N(v_1), N(v_2), \dots, N(v_{k-1})\}}{\operatorname{argmax}} \mathbf{E}[S(\sigma) \mid x_1 = v_1, \dots, x_{k-1} = v_{k-1}, x_k = v_k]$$

end for

9. HW10

1. <https://people.math.osu.edu/nguyen.1261/6501/Note-random1.pdf>

2. Discussed in class. One of the difficulties might be random variables are not independent.

3. A direct applicaiton of LLL. $p = \frac{2}{2^r}, d \leq 2^{r-3}$. $4pd = 4 \frac{2}{2^r} d \leq 4 \frac{2}{2^r} 2^{r-3} = 1$

10. HW11

1. A direct applicaiton of LLL. $p = 2^{1-\binom{k}{2}}, d \leq \binom{k}{2} \binom{n}{k-2}$. $4pd = 4 \binom{k}{2} \binom{n}{k-2} 2^{1-\binom{k}{2}}$

2. A direct applicaiton of LLL. $p = \frac{1}{8^r}, d \leq 2(r-1)$. $4pd = 4 \frac{1}{8^r} 2(r-1) < 1$

3. <https://people.eecs.berkeley.edu/~sinclair/cs271/n24.pdf> Theorem 24.11

REFERENCES

[Erdos and R&WI, 1959] Erdos, P. and R&WI, A. (1959). On random graphs i. *Publ. Math. Debrecen*, 6:290–297.

[Mitzenmacher and Upfal, 2005] Mitzenmacher, M. and Upfal, E. (2005). *Probability and computing: Randomized algorithms and probabilistic analysis*. Cambridge university press.

11. KEY POINTS FOR FINAL

- Chernoff bound technique: Moment-generating function + Markov inequality
- bin and ball model
- Union bound
- 2 forms of local lemma: $4pd$, $x_i \prod_{j \in \Gamma(i)} (1 - x_j)$